







# Simulation-based Bayesian inference for high dimensional inverse problems: Application to magnetic resonance fingerprinting

#### Florence Forbes

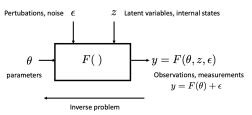
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#### A methodological framework for inverse problems

Common situation: Recover the causes from the observation of the effects



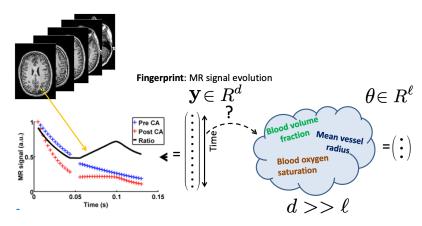
F: analytical formula, mechanistic model, simulator, black box

Inversion:  $F^{-1}$  ? instability, non-unicity or existence of the solution  $\longrightarrow$  III-posed problem

#### Example:

ullet Retrieving brain microvascular properties  $(oldsymbol{ heta})$  from MRI  $(\mathbf{y})$ 

# Retrieving brain microvascular properties from MRI



# [Lemasson et al 2016]

One signal per voxel to invert, d=32,  $\ell=3$ 

#### Bayesian formulation

#### A data generating model

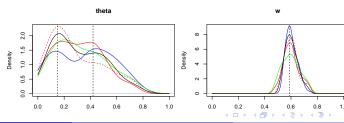
- Prior:  $\theta \sim \pi(\theta)$
- Likelihood:  $f_{\theta}(\mathbf{z})$

$$\longrightarrow \mathbf{z} = \{z_1, \dots, z_d\}$$
 can be simulated from  $f_{m{ heta}}$ 

**Goal:** Estimation of  $\theta$  given some observed  $\mathbf{y} = \{y_1, \dots, y_d\}$ 

Posterior: 
$$\pi(\boldsymbol{\theta}|\mathbf{y}) \propto \pi(\boldsymbol{\theta}) f_{\boldsymbol{\theta}}(\mathbf{y})$$

Prediction:  $\hat{\boldsymbol{\theta}}(\mathbf{y}) = \arg\max_{\boldsymbol{\theta}} \pi(\boldsymbol{\theta}|\mathbf{y})$  or  $\hat{\boldsymbol{\theta}}(\mathbf{y}) = \mathbb{E}[\boldsymbol{\Theta}|\mathbf{y}]$  + uncertainty quantification (eg. variance, multiple modes)



# Problem characteristics (commonly encountered in applications)

- Modelisation of the phenomenon (F), or direct/forward model
  - Theoretical F: imperfect, approximation  $\Rightarrow \mathbf{y} = F(\theta) + \epsilon$  e.g.  $f_{\theta} = \mathcal{N}(F(\theta), \Sigma_{\epsilon})$  (noise known)
  - Complex F:  $\pi(\theta|\mathbf{y})$  known up to a normalising constant. Ok to compute a max, not ok for uncertainty assessment: use MCMC, Variational inference
  - No formula for F: no formula for  $f_{\theta}$  or too costly to evaluate, just a simulator (deterministic or random). Likelihood only known implicitly via simulations
  - Evaluation or simulation is costly: limit the number of required simulations. Can we simulate online?
- High dimension of y, of  $\theta$
- ullet Recover heta and its reliability, multiple solutions. Bayesian posterior? Or point estimate sufficient?
- Massive inversion, scalability: invert a large number of y, repeated inference
- Parameter domain knowledge: choice of prior, summary stats available or not?

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#### Existing likelihood-free or "simulation-based" approaches

- Inverse regression models (e.g. SIR, GLLiM, invertible Neural Nets, Normalizing Flows)
  - ▶ Pros: high dimensional y ok, massive inversion (amortization) ok
  - ▶ Cons: learning cost vs accuracy ? Theoretical garanties?
- Approximate Bayesian Computation (ABC)
- Bayesian Synthetic Likelihood (BSL)

**Remark:** Optimization, Markov Chain Monte Carlo (MCMC) ... , not "simulation-based", require F or the likelihood

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#### Inverse regression: dealing with high dimensional data

Inverse problem: Find  $f(F^{-1})$  between  $\mathbf{y} \in \mathbb{R}^d$  and  $\boldsymbol{\theta} \in \mathbb{R}^\ell$  with  $d >> \ell$ 

$$f: \quad \mathbf{y} \in \mathbb{R}^d \quad \longrightarrow \quad oldsymbol{ heta} \in \mathbb{R}^\ell \quad ext{(high-to-low)}$$

Regression from  $\mathcal{D}_M = \{(\mathbf{z}_m, \boldsymbol{\theta}_m), m = 1 : M\} : d \text{ large} \Longrightarrow \text{curse of dimensionality}$ 

• Inverse regression: learn the low-to-high mapping (F)

Approximate F by  $\hat{F}_M$  by learning a regression model on  $\mathcal{D}_M$  (put the effort here, offline)

Deduce  $F^{-1}$  as  $\hat{F}_M^{-1}$  (use an invertible model in the first place) instead of learning  $F^{-1}$  directly

- Set  $\hat{\theta}(\mathbf{y}) = \hat{F}_M^{-1}(\mathbf{y})$  (straightforward,  $\hat{\theta}(\mathbf{y})$  may not be in  $\mathcal{D}_M$ , no distance computation)
- ullet Add uncertainty modelling by learning the likelihood/posterior instead of  $F/F^{-1}$

GLLiM [Deleforge et al 2015a]:  $f_{\theta}$  and  $\pi(\theta|\mathbf{y})$  are approximated by mixtures of Gaussians, efficient for both sampling and likelihood evaluation.

# Gaussian Locally-linear Mapping (GLLiM)

#### [Deleforge, Forbes & Horaud, 2015a]

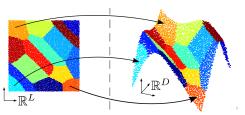
- $oldsymbol{\Theta} \in \mathcal{X} \subset \mathbb{R}^\ell$  low-dimensional space,  $oldsymbol{Y} \in \mathcal{Y} \subset \mathbb{R}^d$  high-dimensional space,
- lacktriangle A piecewise affine model: Introduce a missing variable Z

$$p(\mathbf{y}, \boldsymbol{\theta}; \boldsymbol{\phi}) = \sum_{k=1}^{K} p(\mathbf{y}|\boldsymbol{\theta}, Z = k; \boldsymbol{\phi}) \ p(\boldsymbol{\theta}|Z = k; \boldsymbol{\phi}) \ p(Z = k; \boldsymbol{\phi})$$

 $Z=k\Leftrightarrow \mathbf{Y}$  is the image of  $\mathbf{\Theta}$  by an affine transformation  $au_k$ 

$$oldsymbol{Y} = \sum_{k=1}^K \mathbb{I}(Z=k) (oldsymbol{\mathsf{A}}_k oldsymbol{\Theta} + oldsymbol{b}_k + oldsymbol{E}_k)$$

 ${\mathbb I}$  Indicator function,  ${\bf A}_k$   $d \times \ell$  matrix,  ${\bf b}_k$  d-dim vector  ${\bf E}_k$ : observation noise in  ${\mathbb R}^d$  and reconstruction error, centered, independent on  ${\bf \Theta}$ ,  ${\bf Y}$ , and Z



 $oldsymbol{E}_k$  : centered Gaussian independent on  $oldsymbol{\Theta}$ ,  $oldsymbol{Y}$ , and Z

$$p(\mathbf{y}|\boldsymbol{\theta}, Z = k; \boldsymbol{\phi}) = \mathcal{N}_d(\mathbf{y}; \mathbf{A}_k \boldsymbol{\theta} + \boldsymbol{b}_k, \boldsymbol{\Sigma}_k)$$

ullet Affine transformations are local: mixture of K Gaussians

$$p(\boldsymbol{\theta}|Z=k;\boldsymbol{\phi}) = \mathcal{N}_{\ell}(\boldsymbol{\theta};\boldsymbol{c}_k,\boldsymbol{\Gamma}_k)$$
  
 $p(Z=k;\boldsymbol{\phi}) = \pi_k$ 

• The set of all model parameters is:

$$\phi = \{c_k, \Gamma_k, \pi_k, \mathbf{A}_k, b_k, \Sigma_k, k = 1 : K\}$$

e.g. Mixtures of Linear Experts when  $\Sigma_k$  is diagonal [Xu et al 1995]

Both conditional densities are Gaussian mixtures parameterized by  $\phi$  (or  $\phi^*$  easily deduced from  $\phi$ ),  $\phi$  is estimated using an EM algorithm.

$$\begin{split} p_G(\mathbf{y}|\boldsymbol{\theta};\boldsymbol{\phi}) &= \sum_{k=1}^K \underbrace{\frac{\pi_k \mathcal{N}_\ell(\boldsymbol{\theta}; \boldsymbol{c}_k, \boldsymbol{\Gamma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}_\ell(\boldsymbol{\theta}; \boldsymbol{c}_j, \boldsymbol{\Gamma}_j)}}_{\eta_k(\boldsymbol{\theta})} \mathcal{N}_d(\mathbf{y}; \underbrace{\mathbf{A}_k \boldsymbol{\theta} + \boldsymbol{b}_k}, \boldsymbol{\Sigma}_k) \\ p_G(\boldsymbol{\theta}|\mathbf{y}; \boldsymbol{\phi}^*) &= \sum_{k=1}^K \underbrace{\frac{\pi_k^* \mathcal{N}_d(\mathbf{y}; \boldsymbol{c}_k^*, \boldsymbol{\Gamma}_k^*)}{\sum_{j=1}^K \pi_j^* \mathcal{N}_d(\mathbf{y}; \boldsymbol{c}_j^*, \boldsymbol{\Gamma}_j^*)}}_{\eta_k^*(\mathbf{y})} \mathcal{N}_\ell(\boldsymbol{\theta}; \underbrace{\mathbf{A}_k^* \mathbf{y} + \boldsymbol{b}_k^*}, \boldsymbol{\Sigma}_k^*) \\ & \boldsymbol{c}_k^* = \mathbf{A}_k \boldsymbol{c}_k + \mathbf{b}_k \\ & \boldsymbol{\Gamma}_k^* = \boldsymbol{\Sigma}_k + \mathbf{A}_k \boldsymbol{\Gamma}_k \mathbf{A}_k^\top \\ & \pi_k^* = \pi_k \\ & \mathbf{A}_k^* = \boldsymbol{\Sigma}_k^* \mathbf{A}_k^\top \boldsymbol{\Sigma}_k^{-1} \\ & \mathbf{b}_k^* = \boldsymbol{\Sigma}_k^* (\boldsymbol{\Gamma}_k^{-1} \boldsymbol{c}_k - \mathbf{A}_k^\top \boldsymbol{\Sigma}_k^{-1} \boldsymbol{b}_k) \\ & \boldsymbol{\Sigma}_k^* = (\boldsymbol{\Gamma}_k^{-1} + \mathbf{A}_k^\top \boldsymbol{\Sigma}_k^{-1} \mathbf{A}_k)^{-1} \end{split}$$

Regression functions straightforward via:  $\mathbb{E}_G[Y|\theta;\phi]$  and  $\mathbb{E}_G[\Theta|y;\phi^*]$ 

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# Illustration

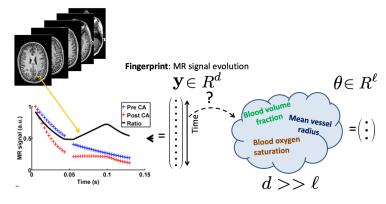
- I. Use GLLiM mixture as an approximation of the posterior
- II. Refine GLLiM approximation with importance sampling
- III. Leverage GLLiM approximation with ABC or BSL

#### I. Use GLLiM to approximate the posterior distributions

- Simulate/Design a dictionary or learning set:  $\mathcal{D}_M = \{(\boldsymbol{\theta}_m, \mathbf{z}_m), m = 1 : M\}$
- ullet Estimate GLLiM parameters (K Gaussians)  $\phi_{K,M}$  from  $\mathcal{D}_M$  using EM and deduce  $\phi_{K,M}^*$
- $\bullet \ \text{Approximate} \ \pi(\boldsymbol{\theta}|\mathbf{y}) \ \text{by} \ p_G(\boldsymbol{\theta}|\mathbf{y};\boldsymbol{\phi}_{K,M}^*) = \textstyle\sum_{k=1}^K \eta_k^*(\mathbf{y}) \ \mathcal{N}_\ell(\boldsymbol{\theta};\mathbf{A}_k^*\mathbf{y} + \mathbf{b}_k^*,\boldsymbol{\Sigma}_k^*)$
- $\bullet$  Predict parameters as  $E_G[\Theta|\mathbf{y};\phi_{K,M}^*] = \sum_{k=1}^K \eta_k^*(\mathbf{y})(\mathbf{A}_k^*\mathbf{y} + \mathbf{b}_k^*)$
- ullet Propose confidence index via  $Var_G[oldsymbol{\Theta}|\mathbf{y};oldsymbol{\phi}_{K,M}^*]$

# Retrieving brain microvascular properties from MRI using MR fingerprinting

[Boux, Forbes, Arbel, Lemasson & Barbier, 2021]

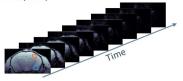


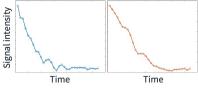
[Lemasson et al 2016]

- High dimensional observations
- Massive inversion, amortization required, MCMC not possible
- Time and memory costly
- Exploitation of the approximated posterior pdf for a confidence index

# Quantitative MRI: Standard MR signal acquisition

> Acquiring images in time according to a specific sequence design (i.e. RF pulses, times, etc.)

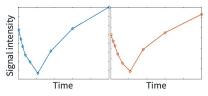




Performing other acquisitions with other designs for an other quantity



One acquisition at a time!



## Standard biophysical parameter quantification

Closed-form expression fitting (CEF): a non-linear least squares solver is used to fit the signals with a biophysical model

> T<sub>2</sub> decay:

$$S_{T_2}(t) = C_{T_2} \exp\left(-\frac{t}{T_2}\right)$$

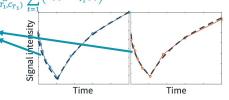
> T<sub>2</sub>\* decay:

$$S_{T_2^*}(t) = C_{T_2^*} \exp\left(-\frac{t}{T_2^*}\right)$$

> T<sub>1</sub> decay:

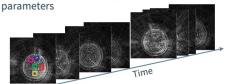
 $S_{T_{2}^{*}}(t) = C_{T_{2}^{*}} \exp\left(-\frac{t}{T_{2}^{*}}\right) \sup_{\substack{s \in S_{T_{1}}(t) \\ s \in S_{T_{1}}(t) \\ s \in S_{T_{1}}(t)}} \left(\frac{S_{T_{2}^{*}}(t) - S_{T_{2}^{*}}(t)}{s}\right) \sup_{\substack{s \in S_{T_{1}}(t) \\ s \in S_{T_{1}}(t)}} \left(\frac{S_{T_{1}}(t) - S_{T_{1}}(t)}{s}\right) \sum_{\substack{s \in S_{T_{1}}(t) \\ s \in S_{T_{1}}(t)}} \left(\frac{S_{T_{1}}(t) - S_{T_{1}}(t)}{s}\right) \exp\left(-\frac{t}{T_{1}}\right)$ 

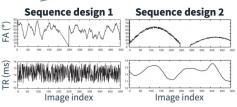
One parameter at a time!



## MR fingerprinting: MR signal (or fingerprint) acquisition in MRF

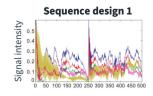
What is really new? Long pseudo-random acquisition sensitive to multiple biophysical

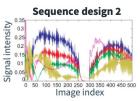




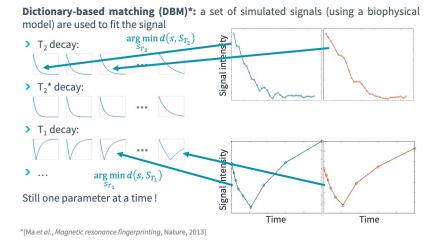
#### Only one acquisition!

[Ma et al., Magnetic resonance fingerprinting, Nature, 2013]





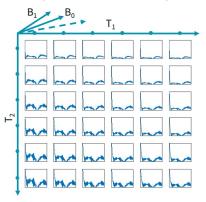
## Standard biophysical parameter quantification in MRF

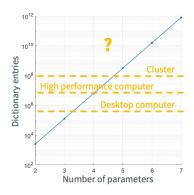


<sup>4</sup> D > 4 A > 4 B > 4 B > B 900

#### Multi-parametric quantification in MRF

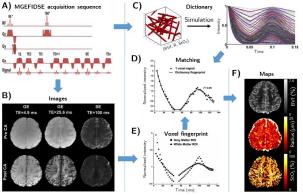
DBM quantification allows multi-parameter quantification from a single fast acquisition





[Ma et al., Magnetic resonance fingerprinting, Nature, 2013]

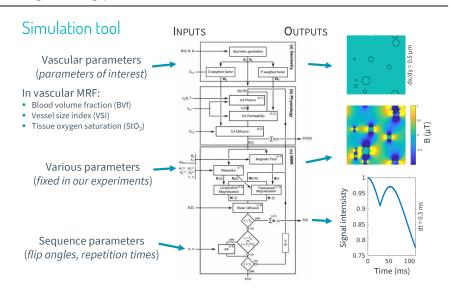
#### Fingerprint: ratio between the MRI signals before and after contrast agent (CA)



[Christen et al., MR vascular fingerprinting: A new approach to compute cerebral blood volume, mean vessel radius, and oxygenation maps in the human brain, Neuroimage, 2014]

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#### Data generating process



#### Limitation: simulation time

	Simulation tool	Ressources	Time		
Lemasson et al.*	Previous tool	30-node cluster	10 <sup>6</sup> signals ≈ 24 hours	2.8 sec/signal	
Pouliot et al.°	Monte Carlo simulations / realistic angiograms	15-node cluster	4.10 <sup>5</sup> signals ≈ 70 hours	10 sec/signal	
This work	Previous tool	32-core computer	10 <sup>5</sup> signals ≈ 67 hours	2.4 sec/signal	
Christen (ANR MRFUSE)	3D previous tool / realistic angiograms	32-core computer	500 signals ≈ 24 hour	3 min/signal	

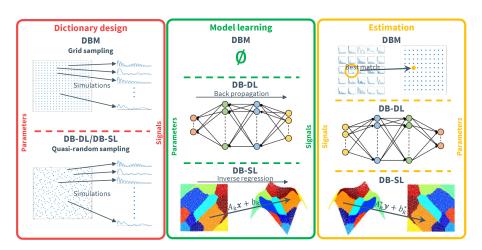
These works suggest that vascular MRF is more accurate than standard vascular approach

#### How to reduce the need for simulations in vascular MRF?

<sup>\*[</sup>Lemasson et al., MR vascular fingerprinting in stroke and brain tumors models, Scientific reports, 2016]

 $<sup>°[</sup>Pouliot\,\textit{et al.}, \textit{Magnetic resonance finger printing based on realistic vasculature in mice}, Neuroimage, 2017]$ 

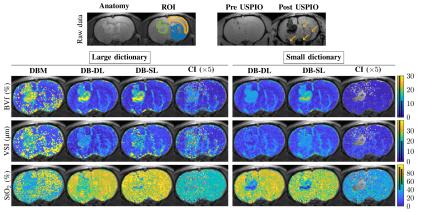
# Analysis framework: comparison



DB-DL: [Cohen et al , MR fingerprinting deep reconstruction network (DRONE), Magnetic Resonance in Medicine 2018]

#### Results: real vascular MRF data

MRI data (4.7T) acquired on a rat bearing a 9L tumor ( $M=167\ 216$  and M=4119)



- bypasses the standard time and memory requirement (Dictionary matching)
- adresses the issue of simulation time
- ullet challenges DL-based quantification (small  $\mathcal{D}_M$ )
- provides confidence maps for each parameter

#### Conclusion and Perspectives

A tractable approach to Bayesian inverse problems, that allows "exploration" of the posterior distribution

#### MR Fingerprinting: a new approach to quantification that

- bypasses the inherent issues in standard MRF (time and memory requirement)
- addresses the problem of simulation time in vascular MRF
- challenges a DL quantification
- provides a confidence map associated to each parameter

#### **Perspectives**

- Real standard MRF acquisitions and complex-valued data
- Spatial considerations (neighboring fingerprints)
- Sequential learning easy with GLLiM
- Sequential learning/Amortization compromise
- Other learning scheme than GLLiM (Mixture density networks, Invertible NN, Normalizing flows) [Lueckmann et al 2021]

# Thank you for your attention!

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#### CODE

- GLLiM, SLLiM and co: xLLiM R package on the CRAN (also GLLiM in Python, Julia)
   Perthame, E., Forbes, F., Deleforge, A., Devijver, E., and Gallopin, M. (2017). xLLiM: High Dimensional Locally-Linear Mapping. R package version 2.1.
- MR fingerprinting (Matlab): https://github.com/nifm-gin/MP3

# Appendix

## Low-to-High or High-to-Low?

If  $\phi$  is unconstrained  $\,\theta$ -to-y or y-to- $\theta$  estimations are equivalent and intractable for large d

Inversion trick: impose a structure on  $\phi$  ( $\phi$  is constrained), e.g.  $\forall k, \Sigma_k = \sigma^2 \mathbf{I}_d$ 

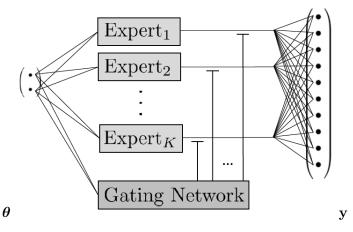
Example: d = 1000,  $\ell = 2$ , K = 10

- Low-to-high regression:  $K(1+\ell+d\ell+\ell(\ell+1)/2+d)=30,060$  parameters.
- High-to-low regression:  $K(1+d+\ell d+d(d+1)/2+\ell)=$  **5,035,030** parameters. Requires inversion of  $1000\times 1000$  covariance matrices.

Therefore it is better to perform a low-dimensional-to-high-dimensional (forward) regression  $(\phi)$ , and then deduce the inverse density  $(\phi^*)$ 

$$\begin{aligned} \boldsymbol{c}_k^* &= \mathbf{A}_k \boldsymbol{c}_k + \mathbf{b}_k \\ \boldsymbol{\Gamma}_k^* &= \boldsymbol{\Sigma}_k + \mathbf{A}_k \boldsymbol{\Gamma}_k \mathbf{A}_k^\top \\ \boldsymbol{\pi}_k^* &= \boldsymbol{\pi}_k \\ \boldsymbol{A}_k^* &= \boldsymbol{\Sigma}_k^* \boldsymbol{A}_k^\top \boldsymbol{\Sigma}_k^{-1} \\ \boldsymbol{b}_k^* &= \boldsymbol{\Sigma}_k^* (\boldsymbol{\Gamma}_k^{-1} \boldsymbol{c}_k - \boldsymbol{A}_k^\top \boldsymbol{\Sigma}_k^{-1} \boldsymbol{b}_k) \\ \boldsymbol{\Sigma}_k^* &= (\boldsymbol{\Gamma}_k^{-1} + \boldsymbol{A}_k^\top \boldsymbol{\Sigma}_k^{-1} \boldsymbol{A}_k)^{-1} \end{aligned}$$

#### GLLiM includes Mixtures of localized linear experts



Expert k :  $p(\mathbf{y}|\boldsymbol{\theta}, Z = k)$  , Gating network:  $p(Z = k|\boldsymbol{\theta})$ 

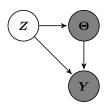
Output:  $p(\mathbf{y}|\boldsymbol{\theta}) = \sum_{k=1}^{K} p(\mathbf{y}|\boldsymbol{\theta}, Z = k) \ p(Z = k|\boldsymbol{\theta})$ 

#### Extension to partially observed responses

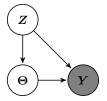
Incorporate a latent component into the low-dimensional variable:

$$\Theta = \left[ egin{array}{c} oldsymbol{T} \ oldsymbol{W} \end{array} 
ight]$$

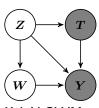
where  $T \in \mathbb{R}^{\ell_{\mathrm{t}}}$  is observed and  $W \in \mathbb{R}^{\ell_{\mathrm{w}}}$  is latent  $(\ell = \ell_{\mathrm{t}} + \ell_{\mathrm{w}})$ 



Supervised GLLiM (regression)



Unsupervised GLLiM (dim reduction)



Hybrid GLLiM

**Remark:** Hybrid GLLiM is supervised GLLiM with  $\Sigma'_k = \Sigma_k + \mathbf{A}_k^w \mathbf{\Gamma}_k^w \mathbf{A}_k^{wT}$  (factor model, e.g. diagonal + low rank matrix)

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#### Particular instances of the hybrid GLLiM model

First three rows: supervised GLLiM methods ( $\ell_{\rm w}=0$ )

Last six rows: unsupervised GLLiM methods ( $\ell_t = 0$ )

Model	$c_k$	$oldsymbol{\Gamma}_k$	$\pi_k$	$\mathbf{A}_k$	$oldsymbol{b}_k$	$oldsymbol{\Sigma}_k$	$\ell_{ m t}$	$\ell_{\mathrm{w}}$	K
MLE [Xu et al 95]	-	-	-	-	-	diag	-	0	-
MLR [Jedidi et al 96]	$0_L$	$\infty$ l $_L$	-	-	-	iso+eq	-	0	-
JGMM [Qiao et al 09]	-	-	-	-	-	-	-	0	-
PPAM [Deleforge et al 12]	-	eq	eq	-	-	diag + eq	-	0	-
GTM [Bishop et al 98]	fixed	$0_L$	eq.	eq.	$0_D$	iso+eq	0	-	-
PPCA [Tipping et al 99a]	$0_L$	$I_L$	-	-	-	iso	0	-	1
MPPCA [Tipping et al 99b]	$0_L$	$I_L$	-	-	-	iso	0	-	-
MFA [Ghahramani et al 96]	$0_L$	$I_L$	-	-	-	diag	0	-	-
PCCA [Bach et al 05]	$0_L$	$I_L$	-	-	-	block	0	-	1
RCA [Kalaitzis et al 11]	$0_L$	$I_L$	-	-	-	fixed	0	-	1

#### Other extensions

- SLLiM: Student mixtures for more robustness: [Perthame et al 2018]
- BLLiM, Structured GLLiM, for specific covariance structures: [Tu et al 2019]
- GLLiM iid: for *i.i.d.* observations: [Forbes et al 2021]
- Markovian GLLiM, for dependent observations: [Deleforge et al 2015c]
- Missing observations: [Deleforge et al 2015b]

