

# Lennard-Jones potential estimation<sup>1</sup>

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Joint work with Frédéric Lavancier



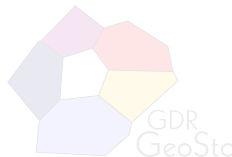
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<sup>1</sup>. based on C-L, Parametric estimation of pairwise Gibbs point processes with infinite range interaction, Bernoulli, 2016



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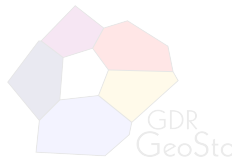
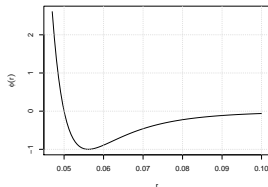




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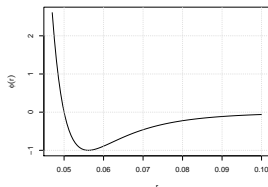




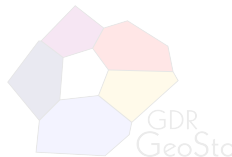
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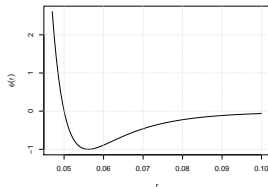




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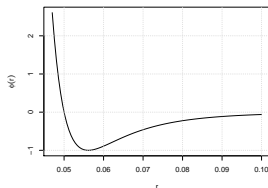




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*Thank you for your attention*



## A few notation

- Let  $\pi_\Lambda^\beta$ ,  $\Lambda \in \mathbb{R}^d$  (bounded Borel set of  $\mathbb{R}^d$ ),  $\beta > 0$  : Poisson point process in  $\Lambda$  with intensity  $\beta$ .
- Let  $\Phi : \mathbb{R}^d \rightarrow \mathbb{R} \cup \{\infty\}$  be a potential function and  $H_\Lambda : \Omega_T \rightarrow \mathbb{R} \cup \{\infty\}$  be the Hamiltonian given by

$$H_\Lambda(\mathbf{x}) = \frac{1}{2} \sum_{\substack{u,v \in \mathbf{x}, u \neq v, \\ \{u,v\} \cap \mathbf{x}_\Lambda \neq \emptyset}} \Phi(u - v)$$

$\Omega = \{\mathbf{x} \in \Omega_T, \forall \Lambda \in \mathbb{R}^d, H_\Lambda(\mathbf{X}) < \infty\}$ ,  $\Omega_T$  set of tempered configurations.

## Dobrushin-Landford-Ruelle formalism

$P$  is a Gibbs measure if  $P(\Omega) = 1$  and for  $P$ -a.e.  $\mathbf{x}$  and any  $\Lambda$  the cond. law of  $P$  given  $x_{\Lambda^c}$  is absolutely continuous w.r.t.  $\pi_\Lambda^\beta$  with density

$$\exp(-H_\Lambda(\mathbf{x})) / Z_\Lambda(\mathbf{x}_{\Lambda^c}).$$



## Papangelou conditional intensity

Let  $\lambda : \mathbb{R}^d \times \Omega \rightarrow \mathbb{R}_+$  defined for any  $u \in \Lambda$  by

$$\lambda(u, \mathbf{x}) = \beta \frac{e^{-H_\Lambda(\mathbf{x} \cup u)}}{e^{-H_\Lambda(\mathbf{x})}} = \beta e^{-\sum_{v \in \mathbf{x}} \Phi(v-u)}.$$

= can be viewed as the conditional probability to have a point in a vicinity of  $u$ , given that the rest of the configuration is  $\mathbf{x}$ .





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Statistical model :  $\Phi \rightarrow \Phi_\theta$ ,  $\lambda \rightarrow \lambda_\theta$ , where  $\theta \in \mathbb{R}^p$ ,  $p \geq 1$

Exponential family model :

$$\lambda_\theta(u, \mathbf{x}) = \beta e^{-\sum_{v \in \mathbf{x}} \Phi_\theta(v-u)} = e^{-\theta^\top t(u, \mathbf{x})}$$

with  $\theta_1 = -\log \beta$  and  $t = (t_1, \dots, t_p)^\top$  where  $t_1(u, \mathbf{x}) = 1$  and

$$t_m(u, \mathbf{x}) = \sum_{v \in \mathbf{x}} g_m(v-u), \quad m = 2, \dots, p.$$

$$\Rightarrow \Phi_\theta = \sum_{m=2}^p \theta_m g_m.$$



## Existence of a Gibbs measure, Ruelle (1969)

If the potential  $\Phi_\theta$  is bounded from below and there exist  $0 < r_1 < r_2 < \infty$ ,  $c > 0$  and  $\gamma_1, \gamma_2 > d$  such that

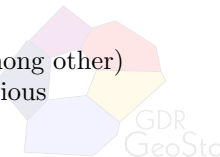
$\Phi_\theta(u) \geq c\|u\|^{-\gamma_1}$  for  $\|u\| \leq r_1$  and  $|\Phi_\theta(u)| \leq c\|u\|^{-\gamma_2}$  for  $\|u\| \geq r_2$ ,  
then  $\exists$  at least one stationary Gibbs measure.

## Examples

- Hard-core with finite range potential!!
- $\Phi(u) = \|u\|^{-\gamma}$  with  $\gamma > d$ ;  $\Phi(u) = e^{-\|u\|}\|u\|^{-\gamma}$  with  $\gamma > d$ .
- Lennard-Jones type pair potential defined for some  $d < \gamma_2 < \gamma_1$  and some  $A, B > 0$  by  $\Phi(u) = A\|u\|^{-\gamma_1} - B\|u\|^{-\gamma_2}$ .
- Standard LJ-model =  $d = 2$ ,  $\gamma_1 = 12$  and  $\gamma_2 = 6$ .

We assume [g] in the following

a collection of technical conditions on the  $g_m$ 's which (among other) imply that  $\Phi_\theta$  satisfies Ruelle conditions. (stified by previous examples)



## Domain of observation

$\mathbf{X}$  is observed in  $W_n$ , where  $(W_n)$  is an increasing sequence of convex domains, such that  $W_n \rightarrow \mathbb{R}^d$  as  $n \rightarrow \infty$ .

## PL and LR

Popular alternatives to the ML (no normalizing constant). In particular, the PL estimator is defined as the maximum of

$$\text{LPL}_{W_n}(\mathbf{X}; \theta) = \sum_{u \in \mathbf{X} \cap W_n} \log \lambda_\theta(u, \mathbf{X} \setminus u) - \int_{W_n} \lambda_\theta(u, \mathbf{X}) du.$$

Problem :  $\lambda_\theta(u, \mathbf{X})$  depends on  $\mathbf{X} \cap W_n^c$  which is not observed.



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**Border-correction** (Møller and Jensen (91), Jensen and Künsch (94), Billiot et al. (08), C and Drouilhet (10), Baddeley et al (14))

$\mathbf{X}$  has a finite range  $\Leftrightarrow \Phi_\theta$  is compactly supported in  $B(0, R)$ ,  $R < \infty$

$\Leftrightarrow \lambda_\theta(u, \mathbf{x}) = \lambda_\theta(u, \mathbf{x}_{u,R})$ , where  $\mathbf{x}_{u,R} = \mathbf{x} \cap B(u, R)$

$$\Rightarrow \text{LPL}_{W_n \ominus R}(\mathbf{X}; \theta) = \sum_{u \in \mathbf{X} \cap W_n \ominus R} \log \lambda_\theta(u, \mathbf{X}_{u,R} \setminus u) - \int_{W_n \ominus R} \lambda_\theta(u, \mathbf{X}_{u,R}) du$$

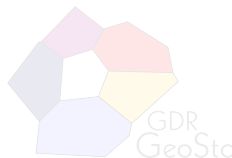


## Pseudolikelihood for infinite range models

Let  $(\alpha_n)$  and  $(R_n)$  be two sequences of real numbers. We define

$$\widetilde{\text{LPL}}_{W_n \ominus \alpha_n, R_n}(\mathbf{X}; \theta) = \sum_{u \in \mathbf{X} \cap W_n \ominus \alpha_n} \log \lambda_\theta(u, \mathbf{X}_{u, R_n} \setminus u) - \int_{W_n \ominus \alpha_n} \lambda_\theta(u, \mathbf{X}_{u, R_n}) du$$

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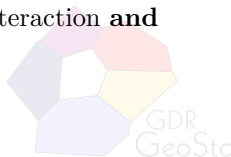
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### Interesting choices (further investigated)

- $\alpha_n = R_n \Leftrightarrow$  border correction for finite range interaction models with range  $R$  taking  $R_n = R$ ;
- $R_n = \infty \Leftrightarrow$  accounting for the maximal possible range of interaction;
- $R_n = \infty$  and  $\alpha_n = 0 \Leftrightarrow$  maximal possible range of interaction **and** no erosion is considered.



## Consistency result

If  $\alpha_n |W_n|^{-1/d} \rightarrow 0$  and  $R_n \rightarrow \infty$  as  $n \rightarrow \infty$  and if for any  $y \in \mathbb{R} \setminus \{0\}$ ,  $P \{y^\top t(0, \mathbf{X}) \neq 0\} > 0$ , then as  $n \rightarrow \infty$ ,  $\widehat{\theta} \rightarrow \theta^*$  a.s.



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## Asymptotic normality result

Let  $\alpha_n$  and  $R_n$  be such that  $\alpha_n |W_n|^{-1/d} \rightarrow 0$ ,  $\alpha_n^{-\gamma'} |W_n|^{1/2} \rightarrow 0$  and  $R_n^{-\gamma'} |W_n|^{1/2} \rightarrow 0$  for some  $0 < \gamma' < \gamma_2 - d$ .

Assume  $U_\infty, \Sigma_\infty > 0$  and  $\boxed{\gamma_2 > 2d}$ , then

$$|W_n|^{1/2} (\widehat{\theta} - \theta^\star) \xrightarrow{d} \mathcal{N}(0, U_\infty^{-1} \Sigma_\infty U_\infty^{-1})$$

- $U_\infty$  : asymptotic normalized sensitivity matrix.
- $\Sigma_\infty$  : asymptotic covariance matrix of the score function.





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## Remarks

- $\gamma_2 > 2d$  is the more important condition (satisfied for LJ).
- for the asymptotic normality result,  $\alpha_n$  must tend to  $\infty$ .



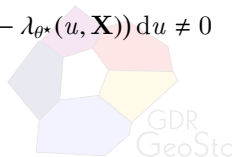
## A few explanations of the difficulties : consistency

(for ease, assume  $\alpha_n = R_n$ )

- for finite range models,  $\text{LPL}_{W_n \ominus R}(\mathbf{X}, \cdot)$  is a concave function and the score function is unbiased.
- for infinite range models,  $\widetilde{\text{LPL}}_{W_n \ominus R_n, R_n}(\mathbf{X}, \cdot)$  can still be shown to be a concave function.
- But its score (at  $\theta^*$ ) writes

$$s_{W_n \ominus R_n}(\mathbf{X}, \theta^*) = \int_{W_n \ominus R_n} t(u, \mathbf{X}_{u, R_n}) \lambda_{\theta^*}(u, \mathbf{X}_{u, R_n}) du - \sum_{u \in \mathbf{X} \cap W_n \ominus R_n} t(u, \mathbf{X}_{u, R_n})$$

$$\Rightarrow \text{E} s_{W_n \ominus R_n}(\mathbf{X}, \theta^*) \stackrel{\text{GNZ}}{=} \int_{W_n \ominus R_n} t(u, \mathbf{X}_{u, R_n}) (\lambda_{\theta^*}(u, \mathbf{X}_{u, R_n}) - \lambda_{\theta^*}(u, \mathbf{X})) du \neq 0$$



## A few explanations of the difficulties : CLT

- CLT for  $\widehat{\theta}$  (mainly) ensues from a CLT for the score function.
- Assume  $W_n \ominus R_n = \cup_{j \in \mathcal{J}_n} \Delta_j$  where the  $\Delta_j$ 's are unit cubes centered at  $j \in \mathbb{Z}^d$ .
- Let  $Z_{n,j} = s_{\Delta_j}(\mathbf{X}, \theta^*) - E(s_{\Delta_j}(\mathbf{X}, \theta^*))$

$$s_{W_n \ominus R_n}(\mathbf{X}, \theta^*) = \sum_{j \in \mathcal{J}_n} s_{\Delta_j}(\mathbf{X}, \theta^*) = \sum_{j \in \mathcal{J}_n} Z_{n,j} + \underbrace{\sum_{j \in \mathcal{J}_n} E s_{\Delta_j}(\mathbf{X}, \theta^*)}_{=o_P(|W_n|^{-1/2})}$$

- The  $Z_{n,j}$ 's are not  $\alpha$ -mixing!! For finite-range models, typically, they are conditionally centered, i.e.

$$E(Z_{n,j} \mid \mathbf{X} \cap \Delta_j^c) = 0 \quad [\text{conditional GNZ}]$$

$\Rightarrow$  CLT for sums of conditionally centered functionals of Markov random fields.



**Theorem** (Jensen and Künsch (1994), Comets and Janzura (1998), Guyon and Gaetan (2004), Coeurjolly and Lavancier (2012))

$Z_{n,j}$  be a triangular array field in a measurable space  $S$ ; For  $n \in \mathbb{N}$ , let  $\mathcal{I}_n \subset \mathbb{Z}^d$  and  $\alpha_n \in \mathbb{R}_+$  such that  $|\mathcal{I}_n| \rightarrow \infty$  and  $\alpha_n \rightarrow \infty$  as  $n \rightarrow \infty$ . Define  $S_n = \sum_{j \in \mathcal{I}_n} Z_{n,j}$  where  $Z_{n,j} = f_{n,j}(X_{n,k}, k \in K_j)$  with  $K_j = \{k \in \mathbb{Z}^d, |k - j| \leq A\}$  and where  $f_{n,j} : S^{K_{n,j}} \rightarrow \mathbb{R}^p$  is a measurable function. We define  $\widehat{\Sigma}_n$  and  $\Sigma_n$  by

$$\widehat{\Sigma}_n = \sum_{j \in \mathcal{I}_n} \sum_{\substack{k \in \mathcal{I}_n \\ |k-j| \leq A}} Z_{n,j} Z_{n,k}^\top \quad \text{and} \quad \Sigma_n = \mathbb{E} \widehat{\Sigma}_n.$$

(a)  $\mathbb{E} Z_{n,j} = 0$  and  $\sup_{n \geq 1} \sup_{j \in \mathcal{I}_n} \mathbb{E} \|Z_{n,j}\|^4 < \infty$ .

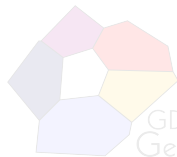
Then,  $|\mathcal{I}_n|^{-1}(\widehat{\Sigma}_n - \Sigma_n) \rightarrow 0$  in  $L^2$ . If in addition

(c)  $\exists Q > 0$  such that  $|\mathcal{I}_n|^{-1} \Sigma_n \geq Q$  for  $n$  sufficiently large,

(d)  $\mathbb{E}(Z_{n,j} \mid X_{n,k}, k \neq j) = 0$

then

$$\Sigma_n^{-1/2} S_n \xrightarrow{d} \mathcal{N}(0, I_p). \quad (1)$$



**Theorem** (Coeurjolly and Lavancier (2016))

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$$\widehat{\Sigma}_n = \sum_{j \in \mathcal{I}_n} \sum_{\substack{k \in \mathcal{I}_n \\ |k-j| \leq \alpha_n}} Z_{n,j} Z_{n,k}^\top \quad \text{and} \quad \Sigma_n = \mathbb{E} \widehat{\Sigma}_n.$$

- (a)  $\mathbb{E} Z_{n,j} = 0$  and there exists  $q \geq 1$  such that  $\sup_{n \geq 1} \sup_{j \in \mathcal{I}_n} \mathbb{E} \|Z_{n,j}\|^{4q} < \infty$ ,  
 (b) for any sequence  $\mathcal{J}_n \subset \mathcal{I}_n$  such that  $|\mathcal{J}_n| \rightarrow \infty$  as  $n \rightarrow \infty$ ,  
 $|\mathcal{J}_n|^{-1} \sum_{j,k \in \mathcal{J}_n} \left\| \mathbb{E} (Z_{n,j} Z_{n,k}^\top) \right\| = O(1)$ .

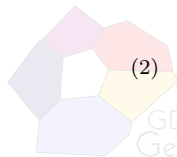
Then if  $\alpha_n^{\frac{4q-1}{2q-1}d} = o(|\mathcal{I}_n|)$  as  $n \rightarrow \infty$ ,  $|\mathcal{I}_n|^{-1} (\widehat{\Sigma}_n - \Sigma_n) \rightarrow 0$  in  $L^{2q}$ . If in addition

- (c)  $\exists Q > 0$  such that  $|\mathcal{I}_n|^{-1} \Sigma_n \geq Q$  for  $n$  sufficiently large,  
 (d) as  $n \rightarrow \infty$

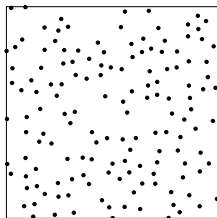
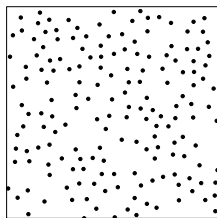
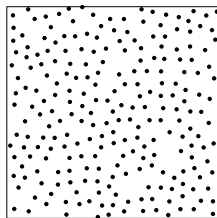
$$|\mathcal{I}_n|^{-1/2} \sum_{j \in \mathcal{I}_n} \mathbb{E} \left\| \mathbb{E} (Z_{n,j} | X_{n,k}, k \neq j) \right\| \rightarrow 0,$$

then

$$\Sigma_n^{-1/2} S_n \xrightarrow{d} \mathcal{N}(0, I_p). \tag{2}$$



- LJ model with parameters :  $\beta = 100$ ,  $\sigma = 0.1$  and  $\varepsilon = 0.1, 0.5, 1$  (low, moderate and high rigidity models).
- $W_n = [-1/2, 1/2]^2, [-1, 1]^2$  and  $[-2, 2]^2$ .

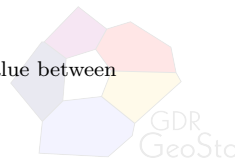
 $\varepsilon = 0.1$  $\varepsilon = 0.5$  $\varepsilon = 1$ 

- $\text{RWMSE} = \sqrt{\text{WMSE}}$  = root-weighted empirical mean squared error, where

$$\text{WMSE} = \frac{\widehat{E} \left\{ \left( \widehat{\log \beta} - \log \beta \right)^2 \right\}}{(\log \beta)^2} + \frac{\widehat{E} \left\{ \left( \widehat{\sigma} - \sigma \right)^2 \right\}}{\sigma^2} + \frac{\widehat{E} \left\{ \left( \widehat{\varepsilon} - \varepsilon \right)^2 \right\}}{\varepsilon^2}$$

	RWMSE		
	$[-1/2, 1/2]^2$	$[-1, 1]^2$	$[-2, 2]^2$
<b>Low (<math>\varepsilon = 0.1</math>)</b>			
$\alpha_n = R_n \in [0.05, 0.3]$	3.26 (0.13)	1.25 (0.13)	0.62 (0.12)
$\alpha_n \in [0.05, 0.3], R_n = \infty$	3.72 (0.05)	1.79 (0.05)	0.63 (0.06)
$\alpha_n = 0, R_n = \infty$	3.5	1.66	0.69
<b>Moderate (<math>\varepsilon = 0.5</math>)</b>			
$\alpha_n = R_n \in [0.05, 0.3]$	0.65 (0.12)	0.34 (0.14)	0.2 (0.15)
$\alpha_n \in [0.05, 0.3], R_n = \infty$	0.68 (0.05)	0.38 (0.05)	0.19 (0.05)
$\alpha_n = 0, R_n = \infty$	0.59	0.33	0.18
<b>High (<math>\varepsilon = 1</math>)</b>			
$\alpha_n = R_n \in [0.05, 0.3]$	1.04 (0.08)	0.42 (0.16)	0.13 (0.16)
$\alpha_n \in [0.05, 0.3], R_n = \infty$	1.34 (0.05)	0.36 (0.05)	0.16 (0.05)
$\alpha_n = 0, R_n = \infty$	1.23	0.27	0.17

- 100 replications
- $\alpha_n$  and/or  $R_n \in [0.05, 0.3] \Leftrightarrow$  30 values regularly sampled.
- When an interval is considered, we report the optimal  $\alpha_n$  value between brackets.



No further question ? Mine : what's the social event ?



(Level imposed by Aalborg was amazingly high)



## Assumption [g]

For convenience we let  $g_1 = 0$  and we denote by  $g$  the  $p$ -dimensional vector  $g = (0, g_2, \dots, g_p)^\top$ . We make the following assumption on  $g$ .

[g] For all  $m \geq 2$ ,  $g_m$  is bounded from below and there exist  $\gamma_1, \gamma_2 > d$  and  $c_g, r_0 > 0$  such that

- (i)  $\forall \|x\| < r_0$  and  $\forall \theta \in \Theta$ ,  $\theta_2 g_2(x) \geq c_g \|x\|^{-\gamma_1}$
- (ii)  $\forall m \geq 3$ ,  $g_m(x) = o(\|x\|^{-\gamma_1})$  as  $\|x\| \rightarrow 0$
- (iii)  $\forall m \geq 2$  and  $\forall \|x\| \geq r_0$ ,  $|g_m(x)| \leq c \|x\|^{-\gamma_2}$ .

