Two nonparametric strategies for estimating the jump rate of a piecewise-deterministic Markov process

Romain Azaïs (Inria Nancy)
Joint work with Aurélie Muller-Gueudin
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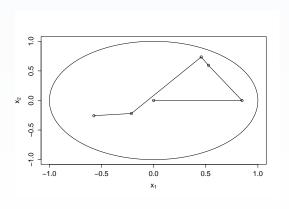


Outline

- Problem formulation
 - Motivation: bacterial motility
 - Piecewise-deterministic Markov process
- 2 Strategies for estimating the jump rate
- 3 Optimal estimation of the jump rate
- 4 Numerical illustration

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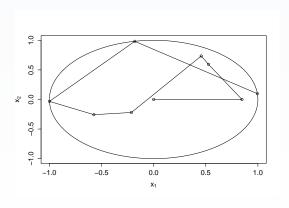
Motivation: bacterial motility







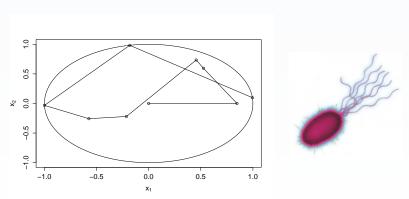
Motivation: bacterial motility







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Is this a "uniform" random walk or is there an attractive chemical signal?



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The survival function of the jump time T_1 is

$$\mathbb{P}(T_1 > t) = \exp\left(-\int_0^t \frac{\lambda}{\lambda} \left(\Phi(x, s)\right) ds\right) \mathbb{1}_{\{\Phi(x, t) \in E\}}.$$



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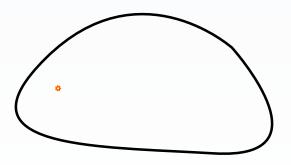
$$\mathbb{P}(T_1 > t) = \exp\left(-\int_0^t \frac{\lambda}{\lambda} \left(\Phi(x, s)\right) ds\right) \mathbb{1}_{\{\Phi(x, t) \in E\}}.$$

At time T_1 the process "jumps" according to Q,

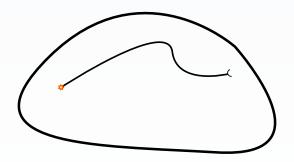
$$\mathbb{E}\left[\varphi(X_{T_1}) \mid \Phi(x, T_1)\right] = \int \varphi(u) \mathcal{Q}(\Phi(x, T_1), du).$$

And so on...

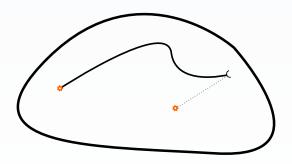




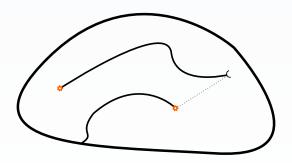




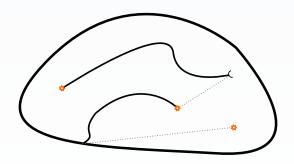




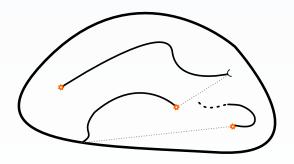




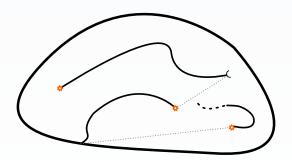












Observation of $Z_n=X_{T_n}$ and $S_{n+1}=T_{n+1}-T_n$ within a long time interval Aim: nonparametric estimation of the jump rate λ



Outline

- Problem formulation
- 2 Strategies for estimating the jump rate
 - Multiplicative intensity model
 - Ratio density function over survival function
- 3 Optimal estimation of the jump rate
- 4 Numerical illustration



Characterization of the jump rate λ :

$$\mathbb{P}(S_{n+1} > t \mid Z_n = x) = \exp\left(-\int_0^t \lambda\left(\Phi(x, s)\right) ds\right) \mathbb{1}_{\{\Phi(x, t) \in E\}}.$$

Conditionnaly on $Z_n=x$, we observe the (right-censored) time S_{n+1} distributed according to the non homogeneous rate $\lambda\circ\Phi(x,t)$.



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Nelson-Aalen strategy

$$M(t) = N(t) - \int_0^t \alpha(s) Y(s) \mathrm{d}s$$
 jump rate of interest



Conditionally on Z_n ,

$$t \mapsto \mathbb{1}_{\{S_{n+1} \le t\}} - \int_0^t \lambda(\Phi(Z_n, s)) \mathbb{1}_{\{S_{n+1} \ge s\}} ds$$

is a continuous-time martingale.

But the sum over n is generally not a martingale



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Solution:

- \rightarrow Estimation for the double-marked renewal process (Z_n, Z_{n+1}, S_{n+1})
- \rightarrow Back to the initial process:
 - discretization of the state space
 - estimation of 2 other functionals



Ratio density function over survival function

Conditionally on $Z_n=x$, the distribution of S_{n+1} admits a density $f(x,\cdot)$ on the interval

$$(0, \inf\{t > 0 : \Phi(x, t) \in \partial E\}),$$

and

$$f(x,t) = \lambda(\Phi(x,t)) \exp\left(-\int_0^t \lambda(\Phi(x,s)) ds\right)$$



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As a consequence,

$$\lambda \circ \Phi(x,t) = \frac{f(x,t)}{G(x,t)}$$

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As a consequence,

$$\lambda \circ \Phi(x,t) = \frac{f(x,t)}{G(x,t)} = \frac{\nu_{\infty}(x)f(x,t)}{\nu_{\infty}(x)G(x,t)}$$



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- Problem formulation
- 2 Strategies for estimating the jump rate
- 3 Optimal estimation of the jump rate
 - Class of estimators indexed by the flow
 - How to choose among this class?
- 4 Numerical illustration

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Class of estimators indexed by the flow

$$\widehat{\mathcal{F}}^{n}(x,t) = \frac{1}{n} \sum_{i=0}^{n-1} \frac{1}{v_i^d w_i} \mathbb{K}_d \left(\frac{Z_i - x}{v_i} \right) \mathbb{K}_1 \left(\frac{S_{i+1} - t}{w_i} \right)$$

$$\widehat{\mathcal{G}}^{n}(x,t) = \frac{1}{n} \sum_{i=0}^{n-1} \frac{1}{v_i^d} \mathbb{K}_d \left(\frac{Z_i - x}{v_i} \right) \mathbb{1}_{\{S_{i+1} > t\}}$$

Estimation of the composed function $\lambda \circ \Phi$

$$\widehat{\lambda \circ \Phi}^{n}(x,t) = \frac{\widehat{\mathcal{F}}^{n}(x,t)}{\widehat{\mathcal{G}}^{n}(x,t)}$$

Optimal estimation of the jump rate

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Class of estimators indexed by the flow

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Estimation of the composed function $\lambda \circ \Phi$

$$\widehat{\lambda \circ \Phi}^{n}(x,t) = \frac{\widehat{\mathcal{F}}^{n}(x,t)}{\widehat{\mathcal{G}}^{n}(x,t)}$$

Pointwise convergence of
$$\widehat{\lambda \circ \Phi}^n(x,t)$$
 $x \in E$ and $\Phi(x,t) \in E$
$$\widehat{\lambda \circ \Phi}^n(x,t) \xrightarrow{a.s.} \lambda \circ \Phi(x,t)$$

$$n^{\frac{1-\alpha d-\beta}{2}} \left(\widehat{\lambda \circ \Phi}^n(x,t) - \lambda \circ \Phi(x,t) \right) \xrightarrow{d} \mathcal{N} \left(0, \frac{\tau_1^2 \, \tau_d^2 \, \lambda \circ \Phi(x,t)}{(1+\alpha d+\beta)\nu_\infty(x) G(x,t)} \right)$$



Class of estimators indexed by the flow

Estimation of $\lambda(x)$ for some $x \in E$

$$\mathcal{C}_x = \{\Phi(x, -t) : t \ge 0\} \cap E$$



For any $\xi \in \mathcal{C}_x$, there exists a unique time $t = \tau_x(\xi)$ such that $\Phi(\xi, \tau_x(\xi)) = x$

In particular,
$$\lambda \circ \Phi(\xi, \tau_x(\xi)) = \lambda(x)$$

Thus,
$$\widehat{\lambda \circ \Phi}^n(\xi, \tau_x(\xi))$$
 estimates $\lambda(x)$, for any $\xi \in \mathcal{C}_x$



How to choose among this class?

We propose to choose among this class the estimate with the minimal asymptotic variance in the central limit theorem:

$$\frac{\tau_1^2 \tau_d^2 \lambda \circ \Phi(\xi, \tau_x(\xi))}{(1 + \alpha d + \beta)\nu_\infty(\xi)G(\xi, t)} \propto (\nu_\infty(\xi)G(\xi, \tau_x(\xi)))^{-1} = \kappa_x(\xi)^{-1}$$

$$\widehat{\lambda}^n(x) = \widehat{\lambda \circ \Phi}^n(\xi^*, \tau_x(\xi^*)) \qquad \text{for } \xi^* = \operatorname*{arg\,max}_{\xi \in \mathcal{C}_x} \kappa_x(\xi)$$

$$\widehat{\widehat{\lambda}}^n(x) = \widehat{\lambda \circ \Phi}^n(\xi^{\star}, \tau_x(\xi^{\star})) \qquad \text{for } \xi^{\star} = \underset{\xi \in \mathcal{C}_x}{\arg \max} \widehat{\mathcal{G}}^n(\xi, \tau_x(\xi))$$



Outline

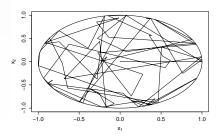
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 - Back to bacterial motility
 - Estimation of the jump rate



Back to bacterial motility

- State space: $E = \{x \in \mathbb{R}^2 : |x| < 1\} \times (0, 2\pi)$
- $\Phi((x_1, x_2, \theta), t) = (x_1 + t \cos \theta, x_2 + t \sin \theta, \theta)$
- $\mathcal{Q}((x_1, x_2, \theta), \cdot) = \delta_{\{x_1, x_2\}} \mathcal{U}_{(0, 2\pi)}$
- $\lambda(x_1, x_2, \theta) = \lambda(x_1, x_2)$

The rate λ is not direction-dependent. But is it really location-dependent?





Estimation of
$$\lambda$$
 at $x = (x_1, x_2, \theta) \in E$

ullet The curve \mathcal{C}_x is given by

$$C_x = \{(x_1 - t\cos\theta, x_2 - t\sin\theta, \theta) : t \ge 0\} \cap E$$

- Compute the optimal point $\xi^* \in \mathcal{C}_x$ that maximizes $\widehat{\mathcal{G}}^n(\xi, \tau_x(\xi))$
- $\bullet \ \widehat{\lambda}^n_{\mathcal{E}^{\star}}(x) \text{ is a good estimate of } \lambda(x) \\$



Estimation of
$$\lambda$$
 at $x = (x_1, x_2, \theta) \in E$

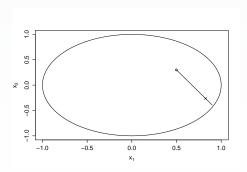
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- Compute the optimal point $\xi^* \in \mathcal{C}_x$ that maximizes $\widehat{\mathcal{G}}^n(\xi, \tau_x(\xi))$
- $\widehat{\lambda}_{\mathcal{E}^{\star}}^{n}(x)$ is a good estimate of $\lambda(x) = \lambda(x_{1}, x_{2}, \theta)$
- By only changing θ , one obtains another good estimate of $\lambda(x_1, x_2)$



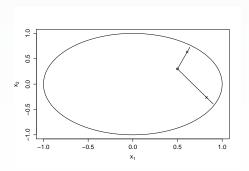
Estimation at
$$x = (0.5, 0.3)$$



$$\begin{array}{c|c} \theta & \widehat{\lambda}_{\xi^*}^n(x) \\ \hline 2\pi/3 & 1.31 \end{array}$$



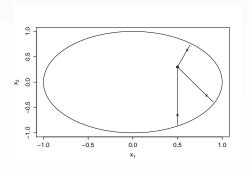
Estimation at
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$$\begin{array}{c|c} \theta & \widehat{\lambda}_{\xi^*}^n(x) \\ \hline 2\pi/3 & 1.31 \\ 7\pi/5 & 0.80 \end{array}$$



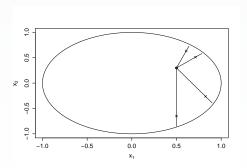
Estimation at x = (0.5, 0.3)



θ	$\widehat{\lambda}_{\xi^{\star}}^{n}(x)$
$2\pi/3$	1.31
$7\pi/5$	0.80
$5\pi/2$	1.41



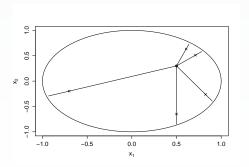
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$5\pi/2$	1.41
$\pi/2$	1.13



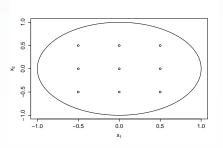
Estimation at x = (0.5, 0.3)



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$2\pi/3$	1.31
$7\pi/5$	0.80
$5\pi/2$	1.41
$\pi/2$	1.13
$\pi/8$	0.91

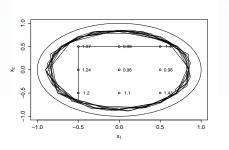


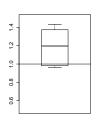
Estimation at 9 target points





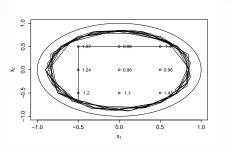
Estimation at 9 target points from $n=\ 20\,000$ observed jumps

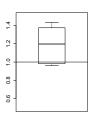






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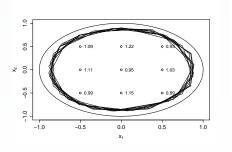


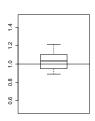


Remark: $\kappa_x(\xi) = \nu_\infty(\xi) G(\xi, \tau_x(\xi)) = \nu_\infty(\xi) \exp(-|x-\xi|)$



Estimation at 9 target points from $n=\ 50\,000$ observed jumps

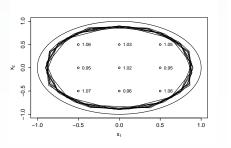


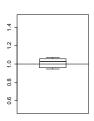


Remark: $\kappa_x(\xi) = \nu_\infty(\xi) G(\xi, \tau_x(\xi)) = \nu_\infty(\xi) \exp(-|x-\xi|)$



Estimation at 9 target points from $n=100\,000$ observed jumps





Remark: $\kappa_x(\xi) = \nu_\infty(\xi) G(\xi, \tau_x(\xi)) = \nu_\infty(\xi) \exp(-|x-\xi|)$